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# Diagnostic Yield of Newborn Screening for Biliary Atresia Using Direct or Conjugated Bilirubin Measurements

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**IMPORTANCE** Treating biliary atresia in newborns earlier can delay or prevent the need for liver transplant; however, treatment typically occurs later because biliary atresia is difficult to detect during its early stages.

**OBJECTIVE** To determine the diagnostic yield of newborn screening for biliary atresia with direct or conjugated bilirubin measurements and to evaluate the association of screening implementation with clinical outcomes.

**DESIGN, SETTING, AND PARTICIPANTS** A cross-sectional screening study of 124 385 infants born at 14 Texas hospitals between January 2015 and June 2018; and a pre-post study of 43 infants who underwent the Kasai portoenterostomy as treatment for biliary atresia at the region's largest pediatric hepatology center before (January 2008-June 2011) or after (January 2015-June 2018) screening implementation. Final follow-up occurred on July 15, 2019.

**EXPOSURES** Two-stage screening with direct or conjugated bilirubin measurements. In stage 1, all newborns were tested within the first 60 hours of life, with a positive screening result defined as bilirubin levels exceeding derived 95th percentile reference intervals. In stage 2, infants who had a positive screening result in stage 1 were retested at or before the 2-week well-child visit, with a positive screening result defined as bilirubin levels greater than the stage 1 result or greater than 1 mg/dL.

MAIN OUTCOMES AND MEASURES The primary outcomes of the screening study were sensitivity, specificity, positive predictive value, and negative predictive value based on infants testing positive in both stages. The reference standard was biliary atresia diagnosed at the region's pediatric hepatology centers. The primary outcome of the pre-post study was the age infants underwent the Kasai portoenterostomy for treatment of biliary atresia.

**RESULTS** Of 124 385 newborns in the screening study, 49.2% were female, 87.6% were of term gestational age, 70.0% were white, and 48.1% were Hispanic. Screening identified the 7 known infants with biliary atresia with a sensitivity of 100% (95% CI, 56.1%-100.0%), a specificity of 99.9% (95% CI, 99.9%-99.9%), a positive predictive value of 5.9% (95% CI, 2.6%-12.2%), and a negative predictive value of 100.0% (95% CI, 100.0%-100.0%). In the pre-post study, 24 infants were treated before screening implementation and 19 infants were treated after screening implementation (including 6 of 7 from the screening study, 7 from screening at nonstudy hospitals, and 6 from referrals because of clinical symptoms). The age infants underwent the Kasai portoenterostomy was significantly younger after screening was implemented (mean age, 56 days [SD, 19 days] before screening implementation vs 36 days [SD, 22 days] after screening implementation; between-group difference, 19 days [95% CI, 7-32 days]; P = .004).

**CONCLUSIONS AND RELEVANCE** Newborn screening with direct or conjugated bilirubin measurements detected all known infants with biliary atresia in the study population, although the 95% CI around the sensitivity estimate was wide and the study design did not ensure complete ascertainment of false-negative results. Research is needed in larger populations to obtain more precise estimates of diagnostic yield and to better understand the clinical outcomes and cost-effectiveness of this screening approach.

Editorial page 1137

Supplemental content

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critical challenge in pediatric hepatology is timely diagnosis of biliary atresia. Biliary atresia is the leading indication for pediatric liver transplant, affecting 1:8000 to 1:18 000 infants worldwide and progressing to endstage liver disease by 2 years of life. The rapid course of biliary atresia can be slowed with the Kasai portoenterostomy, which is an operation that attempts to establish bile flow by removing atretic bile ducts and creating a liver-intestine anastomosis. Treatment by 30 days of life has the best chance of delaying or preventing the need for transplant. In the United States, however, treatment usually occurs later because biliary atresia is difficult to detect during its early stages.

The American Academy of Pediatrics has recommended more studies to develop strategies to help detect biliary atresia earlier. One strategy, the stool color card program, screens infants for the pale stools that develop with biliary obstruction. Stool color card programs have improved treatment times and transplant-free survival in Japan and Taiwan, but they require infrastructure that is currently unavailable in the United States, such as a national call center to triage parents' concerns and a standard 1-month well-child visit to review stool color. A second potential strategy uses newborn direct or conjugated bilirubin measurements, which are elevated in biliary atresia starting shortly after birth.

This 2-part study included a screening study to determine the diagnostic yield of newborn direct or conjugated bilirubin measurements for biliary atresia and a pre-post study to evaluate the association of screening implementation with clinical outcomes.

## Methods

This study was approved by the institutional review board at Baylor College of Medicine and all the participating study sites. For the screening study, a waiver of informed consent was approved because each site was already measuring total bilirubin (and accompanying direct or conjugated bilirubin) in the first 60 hours of life to assess the need for phototherapy based on guidelines from the American Academy of Pediatrics. <sup>15</sup> In addition, the investigator (S.H.) reviewing the results was a credentialed physician at each hospital, and abnormal screening results were shared with clinicians or parents to potentially benefit the patient. For the pre-post study, a waiver of informed consent was approved because the data were collected retrospectively from electronic medical records.

#### Patient Selection for the Screening Study

Infants born at 14 south Texas hospitals between January 2015 and June 2018 were included (eTable 1 in the Supplement). Each hospital had a preexisting policy of measuring total bilirubin in the first 60 hours of life to assess the need for phototherapy. Direct or conjugated bilirubin levels (depending on the type of assay used at each site) were automatically reported alongside total bilirubin levels, although the significance of the values outside the reference intervals was not known and action was not routinely taken.

### **Key Points**

**Question** What is the diagnostic yield of newborn screening for biliary atresia using direct or conjugated bilirubin measurements?

**Findings** In this study that involved 124 385 newborns, a 2-stage screening approach based on direct or conjugated bilirubin measurements identified the 7 known infants with biliary atresia with a sensitivity of 100.0% and a specificity of 99.9%, although the 95% CI around the sensitivity was wide and the study design did not ensure complete ascertainment of false-negative results.

Meaning These findings may help inform decision-making about newborn screening for biliary atresia, although further research is needed from larger populations to obtain more precise estimates of diagnostic yield and to better understand clinical outcomes and cost-effectiveness of this screening approach.

#### Screening Mechanics for the Screening Study

Screening was based on the observation that infants with biliary atresia have elevated direct or conjugated bilirubin levels starting at birth, and involved implementing a standardized clinical response to levels exceeding the reference intervals. <sup>9-12</sup> The upper limit of normal for direct bilirubin was derived at each site using the 97.5th percentile value from 120 newborns consecutively measured. <sup>16</sup> Bilirubin reference intervals were recalculated yearly with new machines to control for differences that occur with changes in the direct bilirubin assay's reaction time, pH, or temperature. <sup>17</sup> The upper limit of normal for conjugated bilirubin was set at 0.2 mg/dL based on established ranges and the consistency of conjugated bilirubin assays across sites. <sup>17,18</sup>

Screening followed a 2-stage prospective direct or conjugated bilirubin testing strategy. In stage 1, all infants were tested via heel stick within the first 60 hours of life as part of routine clinical assessment for hyperbilirubinemia. <sup>15</sup> The infants who had values within their hospital's reference interval were considered to have negative screening results in stage 1 and were not tested further. The infants who had values exceeding their hospital's reference interval were considered to have positive screening results in stage 1, and underwent repeat testing in stage 2 at or before the 2-week well-child visit as part of standard clinical care. For infants no longer at the hospital, the study team called outpatient clinicians to confirm they knew about the abnormal screening results. In stage 2, infants were considered to have negative screening results if they had repeat direct or conjugated bilirubin levels less than or equal to their stage 1 levels and 1 mg/dL or less, which is the 2-week cutoff recommended by the North American and European pediatric gastroenterology societies. 19 Infants with repeat bilirubin levels that were increasing or were greater than 1 mg/dL were considered to have positive screening results in stage 2. Infants with positive screening results in both stages 1 and 2 underwent further testing at the discretion of their clinical care team with consultation from the study team, specialists, or both.

#### **Outcomes for the Screening Study**

The primary outcomes were prespecified and included the sensitivity, specificity, positive predictive value (PPV), and negative predictive value (NPV) of 2-stage screening. The reference standard was biliary atresia diagnosed by intraoperative cholangiogram and pathological examination of the biliary remnant.19 Infants meeting this reference standard were counted using 2 strategies. First, all infants with positive screening results in stage 1 (including those not undergoing stage 2 testing) were followed up until biliary atresia was diagnosed or excluded. Biliary atresia was excluded if direct or conjugated bilirubin levels had normalized, icterus was absent by 3 months per report from the parents or clinicians, or an alternative diagnosis was made. Second, infants with negative screening results in stage 1 were followed up indirectly by monitoring admissions at 3 pediatric hepatology centers in the region. These centers were the referral centers for the birth hospitals included in this study.

# Patient Selection for the Pre-Post Screening Implementation Study

Infants who underwent the Kasai portoenterostomy by 90 days of life at the region's largest pediatric hepatology center (Texas Children's Hospital) were included. A 90-day cutoff was used because the Kasai portoenterostomy is seldom beneficial after 90 days.20 The prescreening group included patients treated between January 2008 and June 2011, which was before screening was implemented and before the value of newborn direct or conjugated bilirubin measurements for detecting biliary atresia was known.9 The postscreening group included patients treated between January 2015 and June 2018, which was after screening had been introduced to the region and implemented at various birth hospitals. Infants in the postscreening implementation group were detected by (1) the screening study, (2) clinicians who learned about the study and implemented a similar direct or conjugated bilirubin screening approach at nonstudy birth hospitals, or (3) clinical symptoms without screening.

# Outcomes for the Pre-Post Screening Implementation Study

The prespecified primary outcome was the age infants underwent the Kasai portoenterostomy for treatment of biliary atresia. Additional outcomes were metrics used in biliary atresia studies, including time to conjugated bilirubin level normalization, transplant-free survival after undergoing the Kasai portoenterostomy, the proportion of infants treated by 30 days, the proportion normalizing their conjugated bilirubin level by 90 days after treatment, and the proportion surviving transplant-free 1 year after treatment. 20-24 The post hoc analyses included time to specialist referral (defined as time from birth to first encounter with a specialist) and time to hepatology evaluation (defined as time from first encounter with a specialist to the Kasai portoenterostomy). Additional post hoc analyses included studying outcomes in subgroups of infants in the group after screening implementation (those detected by the screening study, by screening outside the study, or by clinical symptoms without screening).

#### **Data Collection**

Data were accrued until July 15, 2019. Data collected for the screening study appear in eTable 2 in the Supplement and those collected for the pre-post study appear in eTable 3 in the Supplement. Data on race and ethnicity were collected to detect correlations with the screening results (screening study) or to detect differences between the comparison groups (prepost study). Race and ethnicity were selected by parents from fixed categories at each hospital and the data were combined across hospitals by the study team.

#### Statistical Analysis

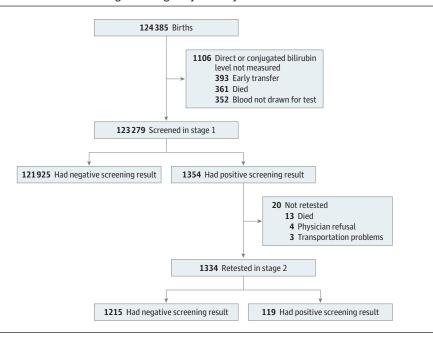
For the screening study, the sample size was calculated using simulations with the bdpv package in R version 2.14.2 (R Foundation for Statistical Computing) to obtain a PPV comparable with newborn screening tests for other rare conditions (0.5%-6.0%). <sup>25</sup> Inputs included a sensitivity of 99%, a specificity of 99%, no assumptions about stage 2 testing, and a North American birth prevalence for biliary atresia of 0.6 per 10 000. <sup>3,26</sup> Based on these assumptions, screening 102 000 newborns was predicted to generate a PPV of 0.67% with a narrow 95% CI of 0.54% to 0.82%. Enrollment stopped at the end of the half-year period in which the calculated sample size was reached.

Birth prevalence calculations included all infants. Diagnostic yield calculations excluded infants not tested in stage 1 and those who had positive screening results in stage 1 but were not retested in stage 2, and used the efficient score method to determine the 95% CIs. Logistic regression was used to model testing status (negative vs positive screening results) in each stage as the dependent variable and patient characteristics (sex, gestational age, race, ethnicity, birth season, first bilirubin assay) as the independent variables. Observations without values were excluded from the respective analyses.

Bivariable analysis was first performed to generate unadjusted 2-sided P values. Multivariable analysis was then performed to adjust for effects of each characteristic. A 2-sided P value of <.008 was considered statistically significant based on a Bonferroni correction for the 6 characteristics evaluated (P = .05/6 = .008) to limit the potential for type I error due to multiple comparisons. Diagnostics on the final models were performed using the Hosmer-Lemeshow goodness-of-fit test.

For the pre-post study, comparisons were made using the unpaired t test (continuous variables), the log-rank test (time to event), and the Fisher exact or  $\chi^2$  test (proportions). For time to normalization of conjugated bilirubin, infants were censored at the time of liver transplant if their conjugated bilirubin level was still elevated. For time to liver transplant, infants who did not receive a transplant were censored on July 15, 2019. Analyses for the pre-post study were 2-tailed, were not adjusted for multiple comparisons, and assumed a P value <.05 indicated statistical significance. For the screening and pre-post studies, all calculations were performed using the VassarStats website, Stata version 14.0 (StataCorp), Prism version 8 (GraphPad), or Excel version 2016 (Microsoft).

Figure 1. Patient Flow for the 2-Stage Screening Study for Biliary Atresia



Stage 1 testing occurred within the first 60 hours of life. Stage 2 testing occurred at or before the 2-week well-child visit

# Results

Of 124 385 newborns in the screening study, 49.2% were female, 87.6% were of term gestational age, 70.0% were white, and 48.1% were Hispanic (eTable 4 in the Supplement). Data were not available on gestational age for 43.6% of infants; race, 12.5%; and ethnicity, 18.1%. Seven infants in the screened population were later diagnosed with biliary atresia (birth prevalence of 0.6 per 10 000).

### Diagnostic Yield of the Screening Study

Stage 1 testing was performed in 99.1% of the study population (n = 123 279; Figure 1). Positive screening results were found in stage 1 testing for 1.1% of newborns (n = 1354). Newborns with positive vs negative screening results in stage 1 were significantly more likely to have been screened with direct bilirubin measurements (91.7% vs 67.3%, respectively; adjusted P < .001), have been born at a preterm gestational age (3.9% vs 1.3%; adjusted P < .001), have been born at an extremely preterm gestational age (2.2% vs 0.7%; adjusted P < .001), and have been identified as black (21.6% vs 13.9%; adjusted P < .001) (Table 1).

Stage 2 testing was performed in 98.5% of infants (n = 1334) who had a positive screening result in stage 1 (Figure 1). Those not retested died before 2 weeks, had clinicians who refused further testing, or lacked transportation to the 2-week well-child visit (eTable 5 in the Supplement). There were 119 infants with positive screening results in stage 2, which comprised 8.9% of those retested and 0.1% of the screened population. Infants with positive vs negative screening results in stage 2 were significantly more likely to be screened initially with conjugated bilirubin measurements (26.1% vs 5.7%, respectively; adjusted P < .001; Table 1).

Stage 1 and 2 results were positive for the 7 known infants with biliary atresia in the study population, resulting in a screening sensitivity of 100.0% (95% CI, 56.1%-100.0%), a specificity of 99.9% (95% CI, 99.9%-99.9%), a PPV of 5.9% (95% CI, 2.6%-12.2%), and a NPV of 100.0% (95% CI, 100.0%-100.0%) (Figure 2). Infants with true positive screening results (n = 7) had initial newborn direct or conjugated bilirubin levels that exceeded site-specific reference intervals to varying degrees in stage 1 and were greater than 1 mg/dL in stage 2 (eTable 6 in the Supplement). Infants with falsepositive screening results (n = 112) tested positive in both stages 1 and 2 but did not have biliary atresia (eTable 7 in the Supplement). Of these 112 infants, 9 (8.0%) had other cholestatic liver diseases, which included Alagille syndrome (n = 4),  $\alpha_1$  antitrypsin deficiency (n = 3), ABCB11 deficiency (n = 1), and choledochal cyst (n = 1) (Table 2). Other falsepositive screening results had no diagnosis determined (52.7%), cholestasis-associated conditions (15.2%), heterozygosity in cholestasis-related genes (10.7%), congenital infections (7.1%), and excessive red blood cell clearance (6.3%). To evaluate the false-positive screening results, additional direct or conjugated bilirubin measurements alone were needed for 25.0% (n = 28) of infants (Table 2), including 37.3% (n = 22) of infants who had no cause of cholestasis determined (eTable 8 in the Supplement).

# Associations With Outcomes During the Pre-Post Screening Implementation Study

Infants who underwent the Kasai portoenterostomy at the region's largest hepatology center before (n = 24) and after (n = 19) screening implementation were compared. The groups had similar demographic factors and similar complication rates after undergoing the Kasai portoenterostomy (eTable 9 in the Supplement). The group after screening implementation

Table 1. Characteristics o	Table 1. Characteristics of Infants Testing Negative or Positive in Each	ositive in Each Screening Stage	ø.					
	Stage 1 (n = $123279$ ) <sup>a,b</sup>				Stage 2 (n = $1334$ ) <sup>a, c,d</sup>			
	No./total (%)				No./total (%)			
Characteristic	Negative screening result	Positive screening result	P value <sup>e</sup>	P value <sup>f</sup>	Negative screening result	Positive screening result	P value <sup>e</sup>	P value <sup>f</sup>
Sex								
Male	61835/121924 (50.7)	732/1354 (54.1)	5	0	653/1215 (53.7)	68/119 (57.1)	Ç	7
Female	60 089/121 924 (49.3)	622/1354 (45.9)	10.	70.	562/1215 (46.3)	51/119 (42.9)	84.	/0:
Gestational age								
Extremely preterm	475/68492 (0.7)	18/826 (2.2)	<.001	<.001	10/730 (1.4)	3/82 (3.7)	.03	.94
Preterm	871/68492 (1.3)	32/826 (3.9)	<.001	<.001	15/730 (2.1)	17/82 (20.7)	<.001	.02
Late preterm	6744/68 492 (9.8)	60/826 (7.3)	.03	.20	42/730 (5.8)	15/82 (18.3)	<.001	.12
Term	60 402/68 492 (88.2)	716/826 (86.7)	Reference	Reference	663/730 (90.8)	47/82 (57.3)	Reference	Reference
Race								
White	74871/106798 (70.1)	738/1099 (67.2)	Reference	Reference	675/993 (68.0)	53/87 (60.9)	Reference	Reference
Black	14847/106798 (13.9)	237/1099 (21.6)	<.001	<.001	207/993 (20.8)	22/87 (25.3)	.26	.83
Asian	7913/106 798 (7.4)	56/1099 (5.1)	.02	.67	53/993 (5.3)	3/87 (3.4)	.59	89.
Other <sup>9</sup>	9167/106 798 (8.6)	68/1099 (6.2)	.03	.05	58/993 (5.8)	9/87 (10.3)	80.	.48
Hispanic	48 122/99 904 (48.2)	512/1061 (48.3)	.95	.95	465/960 (48.4)	40/84 (47.6)	68.	.60
Birth season								
Winter	30006/121925 (24.6)	305/1354 (22.5)	Reference	Reference	272/1215 (22.4)	29/119 (24.4)	Reference	Reference
Spring	34648/121925 (28.4)	385/1354 (28.4)	.25	.04	344/1215 (28.3)	33/119 (27.7)	69.	.93
Summer	28 443/121 925 (23.3)	314/1354 (23.2)	.31	.05	287/1215 (23.6)	22/119 (18.5)	.26	.56
Fall	28828/121925 (23.6)	350/1354 (25.8)	.02	.67	312/1215 (25.7)	35/119 (29.4)	.85	.49
First bilirubin test								
Direct	82 013/121 925 (67.3)	1242/1354 (91.7)	,	,	1146/1215 (94.3)	88/119 (74.0)		,
Conjugated	39912/121925 (32.7)	112/1354 (8.3)	<.001	<.u01	69/1215 (5.7)	31/119 (26.1)	<.uu	<.001

<sup>e</sup> Two-sided test evaluating the hypothesis that selected characteristics are associated with the screening result for both stages in the unadjusted bivariable analyses.

> <sup>o</sup> The Hosmer-Lemeshow goodness-of-fit test was used for the multivariable logistic regression The Hosmer-Lemeshow goodness-of-fit test was used for the multivariable logistic regression

<sup>a</sup> The differences during each stage of screening were modeled using logistic regression.

 $^{\rm d}$  There were 20 infants with positive screening results in stage 1 but were not tested in stage 2.

model(P = .49)model(P = .31).

<sup>&</sup>lt;sup>f</sup> Two-sided test evaluating the hypothesis that selected characteristics are associated with the screening result for both stages in the multivariable analyses adjusted for sex, gestational age, race, ethnicity, birth season, and first bilirubin test.

<sup>&</sup>lt;sup>3</sup> Included American Indian, Alaskan Native, or self-reported as other.

Figure 2. Newborn Direct or Conjugated Bilirubin Screening for Biliary Atresia

Stage 1	Positive screening result	Negative screening result	Total No.
Positive	7	1347	1354
Negative	0	121925	121925
Total	7	123 272	123 279
	Docitivo	Nogativo	
Stage 2	Positive screening result	Negative screening result	Total No.
Stage 2 Positive			Total No. 119
•		screening result	1

Total	Positive screening result	Negative screening result	Total No.
Positive	7	112	119
Negative	0	123 140	123 140
Total	7	123252	123259

% (95% CI)				
Sensitivity	100.0 (56.1-100.0)			
Specificity	99.9 (99.9-99.9)			
PPV	5.9 (2.6-12.2)			
NPV	100.0 (100.0-100.0)			

NPV indicates negative predictive value; PPV, positive predictive value.

consisted of 6 infants detected by the screening study (1 infant with a true positive screening result underwent the Kasai portoenterostomy at another study site) and 7 infants detected by clinicians at nonstudy birth hospitals that had adopted 2-stage clinical intervention programs based on the screening study (eTable 10 in the Supplement). The group after screening implementation also included 6 infants born at hospitals not participating in the study who were referred because of concerning clinical symptoms.

The age infants underwent the Kasai portoenterostomy was significantly younger after screening was implemented (mean age, 56 days [SD, 19 days] before implementation vs 36 days [SD, 22 days] after implementation; between-group difference, 19 days [95% CI, 7-32 days], P = .004; Figure 3A). After screening implementation, infants also were significantly more likely to undergo the Kasai portoenterostomy by 30 days (12.5% before implementation vs 57.9% after implementation; reciprocal of risk ratio, 4.6 [95% CI, 1.7-14.0], P = .003; Table 3).

In the post hoc analyses, infants were referred to a specialist at significantly younger ages after screening implementation (mean age, 44 days [SD, 24 days] before implementation vs 25 [SD, 21] days after implementation; between-group difference, 20 days [95% CI, 6 to 34 days], P = .007; Figure 3B). However, after screening implementation, infants were not evaluated faster after referral (mean time for evaluation, 11 days [SD, 13 days] before implementation vs 12 days [SD, 12 days] after implementation; between-group difference, 0 days [95%

CI, -8 to 7 days], P = .92; Figure 3B). Within the group after screening implementation, infants identified by clinical symptoms were referred and treated later, whereas infants screened either at study hospitals only or at study hospitals and non-study hospitals were more likely to undergo the Kasai portoenterostomy by 30 days (eTables 11 and 12 and eFigure in the Supplement).

For outcomes after undergoing the Kasai portoenterostomy, infants in the group after screening implementation compared with those in the group before screening implementation had significantly faster times to conjugated bilirubin normalization (log-rank P = .03; Figure 3C) but did not have a statistically significant difference in transplant-free survival (log-rank P = .13; Figure 3D). Infants in the group after screening implementation were significantly more likely to have normal conjugated bilirubin levels by 90 days after undergoing the Kasai portoenterostomy (41.7% before implementation vs 78.9% after implementation; reciprocal of risk ratio, 1.9 [95% CI, 1.1-3.3], P = .03; Table 3). Transplantfree survival rates 1 year after undergoing the Kasai portoenterostomy were higher but did not reach statistical significance (70.8% before screening implementation vs 94.7% after screening implementation; reciprocal of risk ratio, 1.3 [95% CI, 1.0-1.9], P = .06; Table 3). Post hoc comparisons showed that infants screened in study and nonstudy hospitals after screening implementation compared with those before screening implementation were more likely to survive transplant-free 1 year after undergoing the Kasai portoenterostomy (eTable 12 in the Supplement).

## Discussion

This study evaluated newborn direct or conjugated bilirubin screening for biliary atresia using 2 approaches. First, in the screening study of a large population of newborns, direct or conjugated bilirubin testing had a sensitivity of 100%, a specificity of 99.9%, a PPV of 5.9%, and a NPV of 100.0%, but with a wide 95% CI for the sensitivity estimate because of the small number of true positive screening results. Second, in the prepost study, screening was associated with earlier treatment; the mean age infants underwent the Kasai portoenterostomy was 56 days before screening implementation vs 36 days after screening implementation. After undergoing the Kasai portoenterostomy, infants in the group after screening implementation had significantly faster times to conjugated bilirubin normalization compared with infants in the group before screening implementation.

The association of screening with younger age when undergoing the Kasai portoenterostomy compares well with other biliary atresia screening methods currently in use. For example, in Taiwan, universal implementation of the stool color card program was associated with a mean age of 46 days when undergoing the Kasai portoenterostomy. In Japan, regional implementation of the stool color card program was associated with a mean age of 59.7 days when undergoing the Kasai portoenterostomy. One additional advantage of newborn direct or conjugated bilirubin screening is that it

<sup>&</sup>lt;sup>a</sup> There were 20 newborns who were not retested in stage 2 because 13 died, the physician refused to test in 4, and there were transportation problems for 3.

Table 2. Diagnoses and Evaluation for False-Positive Screening Results (n = 112)

Description of diagnosis and evaluation	No. (%)
Type of diagnosis	
Not determined	59 (52.7)
Cholestasis-associated conditions <sup>a</sup>	17 (15.2)
Heterozygosity in cholestasis-related genes <sup>b</sup>	12 (10.7)
Cholestatic liver diseases <sup>c</sup>	9 (8.0)
Congenital infections <sup>d</sup>	8 (7.1)
Excessive red blood cell clearance	7 (6.3)
Type of evaluation performed	
Additional direct or conjugated bilirubin testing only	28 (25.0)
Additional laboratory testing	25 (22.3)
Additional noninvasive imaging	38 (33.9)
Liver biopsy with or without percutaneous transhepatic cholangiogram	20 (17.9)
Intraoperative cholangiogram	1 (0.9)

- <sup>a</sup> Included trisomy 21 (5 cases), gastroschisis (4 cases), trisomy 18 (3 cases), portosystemic shunt (2 cases), maternal lupus (1 case), omphalocele (1 case), and panhypopituitarism (1 case).
- <sup>b</sup> The gene names appear in eTable 7 in the Supplement.
- <sup>c</sup> Included Alagille syndrome (4 cases), α<sub>1</sub> antitrypsin deficiency (3 cases), *ABCB11* deficiency (1 case), and choledochal cyst (1 case).
- <sup>d</sup> Included cytomegalovirus (3 cases), syphilis (3 cases), coxsackievirus (1 case), and rubella (1 case).

potentially allows infants to undergo the Kasai portoenterostomy before 30 days, the time interval correlating with the best chance of delaying or preventing the need for transplant. <sup>2,3</sup> In the pre-post study, 57.9% of infants underwent the Kasai portoenterostomy by 30 days after screening implementation compared with 12.5% before screening implementation.

Screening also identified 112 infants without biliary atresia, resulting in a PPV of 5.9% that is comparable with the PPV range of 0.5% to 6.0% reported for other newborn screening tests. <sup>25</sup> Some of the infants with false-positive screening results had conditions that could have benefitted from early detection and intervention. For example, infants with cholestatic liver diseases such as Alagille syndrome and  $\alpha_1$  antitrypsin deficiency have been reported to present with life-threatening bleeding associated with vitamin K deficiency. <sup>27,28</sup> Detection during the newborn period allowed for nutritional supplementation and vitamin K administration as needed. In addition, infants with congenital infections such as cytomegalovirus infection can be identified and treated quickly, which may help limit development of serious neurological sequelae. <sup>29</sup>

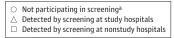
An important avenue of future research is the cost-effectiveness of screening. A recent Canadian study estimated an incremental cost-effectiveness ratio for bilirubin screening to be unacceptable at \$473 840 per life-year gained. The wever, this calculation relied on assumptions not supported by data from this study, including a specificity of 98%, a full hepatology evaluation for all false-positive screening results, and 37% of infants undergoing the Kasai portoenterostomy by 30 days. The study also used life-years gained as a metric for effectiveness, whereas transplant-free life-years gained may also be informative because biliary atresia seldom results in death. Additional simulations using new information (including data from this study) can further assess the cost-effectiveness of screening.

One way to potentially improve cost-effectiveness is by reducing the number of infants who test positive in stage 1 or stage 2 but who are unlikely to have serious liver conditions. In stage 1, more than 90% of positive screening results were negative when retested in stage 2. The infants with positive

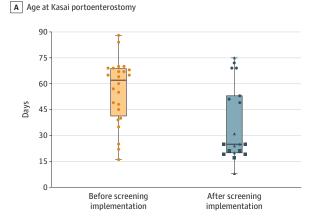
screening results in stage 1 were more likely to have been tested with direct vs conjugated bilirubin, reflecting the inherent limitations of the direct bilirubin assay that also measures small amounts of unconjugated bilirubin.<sup>17</sup> Even though some have suggested replacing the direct assay with the conjugated assay to avoid these limitations, 2 hospitals did the opposite during the study period and switched from conjugated assays to direct assays for cost and convenience reasons. 18 In stage 2, 25% of infants who had positive screening results had subsequent normal direct or conjugated bilirubin levels and did not require further testing. Some of these false-positive results could have been avoided by performing stage 2 testing later; however, for practical reasons, stage 2 testing is best performed at the routine newborn follow-up well-child visit recommended by the American Academy of Pediatrics.

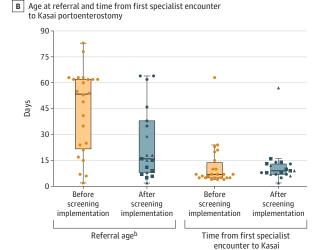
Even if studies with larger populations confirm the results from this study and screening is found to be costeffective, there are still important implementation challenges that need to be addressed. First, birth hospitals would need to carefully derive and update site-specific reference intervals (techniques to screen newborn blood spots at central laboratories have not been developed). For example, 1 infant with a true positive screening result had a stage 1 bilirubin level of 0.4 mg/dL, which was abnormally high in the 98.4th to 99.6th percentile for that site but may have been overlooked if the reference intervals were inappropriately borrowed from another site. Second, stage 2 testing requires reliable access to medical centers willing to participate in screening. In this study, 0.2% of infants with positive screening results in stage 1 had parents or caregivers who could not travel to the well-child visit, and another 0.2% had clinicians who chose not to retest. Third, infants with a positive screening result must be quickly evaluated for biliary atresia, which can be difficult to detect during early stages of the disease when aminotransferase levels and liver stiffness may still be normal.31,32 The challenge specialists face was highlighted by an infant who had a true positive screening result in the study, but who underwent the Kasai portoenterostomy at 75 days despite identification at 2 weeks by screening.

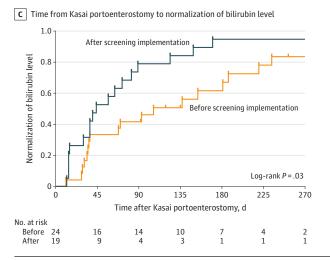
Figure 3. Outcomes Associated With Implementation of Screening for Biliary Atresia

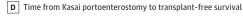


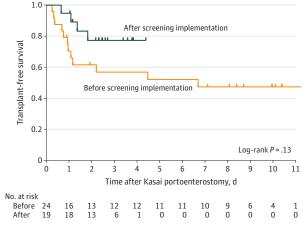
portoenterostomy











In A and B, the box plot lines correspond from the bottom of box to the top: 25th percentile, median percentile, and 75th percentile. The whiskers extend to the upper and lower adjacent values. In C, the median follow-up time was 100 days (quartile 1-3, 36-186 days) in the group before screening implementation and 44 days (quartile 1-3, 15-89 days) in the group after screening implementation. In D, the median follow-up time was 3.3 years (quartile 1-3, 0.9-9.6 years) in the group before screening implementation and

- 2.5 years (quartile 1-3, 1.4-3.6 years) in the group after screening implementation
- <sup>a</sup> The circles that represent infants in the group after screening implementation had biliary atresia detected by clinical symptoms rather than newborn direct or conjugated bilirubin tests.
- <sup>b</sup> Defined as time from birth to first encounter with a specialist.

#### Limitations

The screening study has several limitations. First, the small number of true positive screening results affected the 95% CI for sensitivity, which was wide at 56.1% to 100%, and lower limits may be unacceptable for a screening test. Second, many infants (43.6%) did not have gestational age recorded, preventing a complete analysis of the association between prematurity and the screening results. Third, screening only occurred at south Texas hospitals, potentially limiting the applicability of the study's results to other populations.

Fourth, a patient with biliary atresia could have been missed if an infant tested negative in stage 1, moved away from the region, and was then diagnosed with biliary atresia at a center other than one of the region's 3 pediatric hepatology referral centers monitored during the study. However, normal direct or conjugated bilirubin levels in the newborn period would be very unlikely among those with biliary atresia. 9-12 In addition, the study's birth prevalence for biliary atresia of 0.6 per 10 000 newborns matches the birth prevalence reported in multiple North American and European studies;

Table 3. Secondary Outcomes Associated With Implementation of Screening for Biliary Atresia

	No. (%)			
Outcome	Before screening implementation (n = 24)	After screening implementation (n = 19)	Reciprocal of risk ratio (95% CI)	P value <sup>a</sup>
Underwent the Kasai portoenterostomy by 30 d of life	3 (12.5)	11 (57.9)	4.6 (1.7-14.0)	.003
Normalization of bilirubin level by 90 d after undergoing the Kasai portoenterostomy <sup>b</sup>	10 (41.7)	15 (78.9)	1.9 (1.1-3.3)	.03
Transplant-free survival 1 y after undergoing the Kasai portoenterostomy	17 (70.8)	18 (94.7)	1.3 (1.0-1.9)	.06

<sup>&</sup>lt;sup>a</sup> Calculated using the Fisher exact test.

however, this prevalence is less than the birth prevalence of 1.6 per 10 000 reported in Taiwan.  $^{7,26}$ 

The pre-post study also has important limitations. First, the sample sizes in the groups before and after screening implementation were small. This may have prevented trends in outcomes such as 1-year transplant-free survival from reaching statistical significance. Second, the pre-post study design used historical controls that may have introduced bias because of temporal changes in care unrelated to screening. To limit biases with historical controls, infants in both groups were treated by the same hepatologists and surgeons at the same hospital. Furthermore, historical controls may be well-suited for biliary atresia studies because the types of care and outcomes have not changed significantly over the past 20 years. For example, the mean age infants undergo the Kasai portoenterostomy has remained the

same, as has the proportion of infants who have normal bilirubin levels or survive transplant-free after undergoing the Kasai portoenterostomy. 4,5,20,23,24,33

#### Conclusions

Newborn screening with direct or conjugated bilirubin measurements detected all known infants with biliary atresia in the study population, although the 95% CI around the sensitivity estimate was wide and the study design did not ensure complete ascertainment of false-negative results. Research is needed in larger populations to obtain more precise estimates of diagnostic yield and to better understand the clinical outcomes and cost-effectiveness of this screening approach.

#### ARTICLE INFORMATION

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**Author Contributions:** Drs Harpavat and Shneider had full access to all of the data in the study and take responsibility for the integrity of the data and the accuracy of the data analysis.

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